# **Bandits: Explore-Then-Commit, -greedy, UCB** *ε*

#### **Lucas Janson CS/Stat 184(0): Introduction to Reinforcement Learning Fall 2024**







- Feedback from last lecture
- Recap
- Regret analysis of ETC
- -greedy algorithm *ε*
- Confidence intervals for the arms
- Upper Confidence Bound (UCB) algorithm

#### Feedback from feedback forms

1. Thank you to everyone who filled out the forms! 2.







- Recap
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#### Recap

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- Multi-armed bandits (or MAB or just bandits)
	- •Online learning of a 1-state/1-horizon MDP
	- •Exemplify exploration vs exploitation
	- •Pure greedy & pure exploration achieve linear regret
	- Hoeffding's inequality

#### Recap

- Multi-armed bandits (or MAB or just bandits)
	- •Online learning of a 1-state/1-horizon MDP
	- •Exemplify exploration vs exploitation
	- •Pure greedy & pure exploration achieve linear regret
	- Hoeffding's inequality
- Today: let's do better than linear regret!

#### Notes from last lecture

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1. 
$$
\text{Regret}_{T} = T\mu^{\star} - \sum_{t=0}^{T-1}
$$

2. Recall Regret $_{T} = \Omega(T)$ , i.e., linear regret  $\Rightarrow$  for some  $c > 0$  and  $T_0$ , Regret<sub> $T \ge cT$ </sub>  $\forall T \ge T_0$ 

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#### m last lecture  $\mu_{a_t} =$ *T*−1 ∑  $\overline{t=0}$  $(\mu^{\star} - \mu_{a_t})$ ) Expected regret at time *t given that you chose arm*  $a_t$ Regret<sub>T</sub> *T*  $\rightarrow 0$  $\Rightarrow$  for some  $c > 0$  and  $T_0$ , Regret<sub> $T \geq cT$ </sub>  $\forall T \geq T_0$ 4. Hoeffding inequality: sample mean of  $N$  i.i.d. samples on  $\left[0,1\right]$  satisfies  $w/p 1 - \delta$ ln(2/*δ*) 2*N*

$$
|\hat{\mu} - \mu| \leq \sqrt{2\pi}
$$

#### Explore-Then-Commit (ETC)  $N_{\odot} =$  <u>Number</u> of explorations

- Algorithm hyper parameter  $N_{\rm e}$  <  $T/K$  (we assume  $T >> K$ )
- For  $k = 1, \ldots, K$ : (Exploration phase)
	- Pull arm  $k$   $N_{\bf e}$  times to observe  $\{r_i^{(k)}\}_{i=1}^{N_{\bf e}}$ ∼ *ν<sup>k</sup>* Calculate arm k's empirical mean:  $\hat{\mu}_k =$ ̂ 1 *N*e *N*e ∑ *i*=1  $r_i^{(k)}$ *i*
- For  $t = N_{\rm e}K, ..., (T-1)$ : (Exploitation phase)

Pull the best empirical arm  $a_t = \arg \max_{i \in [V]} \hat{\mu}$ *i*∈[*K*] ̂ *i*



### Regret Analysis Strategy

- 1. Calculate regret during exploration stage
- 2. Quantify error of arm mean estimates at end of exploration stage
- 3. Using step 2, calculate regret during exploitation stage
	- (Actually, will only be able to upper-bound total regret in steps 1-3)
- 4. Minimize our upper-bound over  $N_{\mathbf{e}}$

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	- b) Recall Union/Boole/Bonferroni bound:  $\mathbb{P}(\mathsf{any} \; \mathsf{of} \, A_1, ..., A_K) \leq 1$ *K* ∑  $\mathbb{P}(A_k)$

$$
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*k*=1

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\sqrt{\ln(2/\delta)/2N_{\mathbf{e}}}\big) \ge 1 - \delta_{\mathbb{P}(\forall k, A_1^c, \dots, A_K^c) \ge 1 - \sum_{k=1}^K \mathbb{P}(k)} \frac{1}{\sqrt[k]{\sum_{k=1}^K \sum_{k=1}^K \mathbb{P}(k)}} \frac{1}{\sqrt[k]{\sum_{k=1}^K \sum
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b) Recall Union/Boole/Bonferroni bound:  $\mathbb{P}(\text{any of } A_1, ..., A_K) \leq \sum_{k=1}^K \mathbb{P}(A_k)$   
c)  $\delta \rightarrow \delta/K$ , Union bound with  $A_k = \left\{ |\hat{\mu}_k - \mu_k| > \sqrt{\ln(2K/\delta)/2N_e} \right\}$ , and Hoeffding:





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Denote (apparent) best arm after exploration stage by  $k$  and actual best arm by  $\hat{k}$  and actual best arm by  $k^{\star}$ regret at each step of exploitation phase  $=\mu_{k^\star}-\mu_{\hat{k}}$ 

> $\mu_k$  + ( $\hat{\mu}_{\hat{k}} - \hat{\mu}_{\hat{k}}$ )  $\ddot{\phantom{a}}$  $\ddot{\phantom{a}}$



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= \mu_{k*} + (\hat{\mu}_{k*} - \hat{\mu}_{k*}) - \mu_{\hat{k}} + (\hat{\mu}_{\hat{k}} - \hat{\mu}_{\hat{k}})
$$
  

$$
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\n
$$
= \sqrt{2\ln(2K/\delta)/N_{\mathbf{e}}}
$$

 $=\sqrt{2 \ln(2K/\delta)/N_{\Theta}}$ 

 $\Rightarrow$  total regret during exploitation

$$
\lim_{n} \leq T\sqrt{2\ln(2K/\delta)/N_{\Theta}} \quad \text{w/p } 1 - \delta
$$


4. From steps 1-3: with probability  $1 - \delta$ ,

- 
- $Regret$ <sub> $T$ </sub>  $\leq N$ <sub>e</sub> $K + T\sqrt{2 \ln(2K/\delta)/N_e}$

- 4. From steps 1-3: with probability  $1 \delta$ ,
	- $Regret_\tau \leq N_eK + T\sqrt{2\ln(2K/\delta)/N_e}$
	- What's a choice of  $N_{\rm e}$  that gives sublinear regret?

- 4. From steps 1-3: with probability  $1 \delta$ ,
	- $\text{Regret}_{T} \leq N_{e}K + T_{v} \sqrt{2 \ln(2K/\delta)/N_{e}}$
	- What's a choice of  $N_{\rm e}$  that gives sublinear regret?
	- Any  $N_e$  so that  $N_e \rightarrow \infty$  and  $N_e/T \rightarrow 0$  (e.g.,  $N_e = \sqrt{T}$ )

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optimal  $N_{\text{e}} =$ 

Minimize over  $N_{\mathbf{e}}$ :

$$
\left(\frac{T\sqrt{\ln(2K/\delta)/2}}{K}\right)^{2/3}
$$

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		- Minimize over  $N_{\mathbf{e}}$ :
		- optimal  $N_{\mathbf{e}} =$
	- $\Rightarrow$  Regret<sub>*T*</sub>  $\leq 3T^{2/3}$  $(K \ln(2K/\delta)/2)^{1/3} = o(T)$ (A bit more algebra to plug optimal  $N_e$  into Regret<sub>*T*</sub> equation above)
- 

$$
\left(\frac{T\sqrt{\ln(2K/\delta)/2}}{K}\right)^{2/3}
$$





• Feedback from last lecture



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 $For t = 0, ..., T - 1:$ Sample  $E_t \sim$  Bernoulli $(\varepsilon)$ Initialize  $\hat{\mu}_0 = \cdots = \hat{\mu}_K = 1$ ̂ ̂



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\nSample  $E_t \sim \text{Bernoulli}(\varepsilon)$ 

\nIf  $E_t = 1$ , choose  $a_t \sim \text{Uni}$ 

#### If  $E_t = 1$ , choose  $a_t \sim \text{Uniform}(1, ..., K)$  (pure explore)  $\sim$  Uniform $(1,...,K)$



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\nOutput

\nDescription:



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- at *every* step, do pure greedy w/p  $1 \varepsilon$ , and do pure exploration w/p  $\varepsilon$ 
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- Can also allow *ε* to depend on *t*; should it increase, decrease, or stay flat? The more learned by time *t*, the less exploration needed at/after time *t*
- It turns out that  $\varepsilon$ -greedy with  $\varepsilon$ <sub>r</sub> =  $\vert$   $\vert$   $\vert$   $\vert$   $\vert$  also achieves  $\varepsilon$ -greedy with  $\varepsilon_t =$
- where  $O({}\cdot{})$  hides logarithmic factors ˜  $(\ \cdot\ )$



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-greedy with  $\varepsilon_t = \left(\frac{K \ln(t)}{t}\right)^{1/3}$  also achieves  
Regret<sub>t</sub> =  $\tilde{O}(t^{2/3}K^{1/3})$ ,

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- where  $O({}\cdot{})$  hides logarithmic factors ˜  $(\ \cdot\ )$ 
	- Regret rate (ignoring log factors) is the same as ETC, but holds for all t, not just the full time horizon *T*
	- Nothing in  $\varepsilon$ -greedy (including  $\varepsilon$ <sub>t</sub> above) depends on  $T$ , so don't need to know horizon!







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Sample mean of  $N$  i.i.d. samples on  $[0,1]$  satisfies

First: how to construct confidence intervals? Recall Hoeffding inequality:

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\hat{\mu} - \mu \mid \leq \sqrt{\frac{\ln(2/\delta)}{2N}} \text{ w/p } 1 - \delta
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Worked for ETC b/c exploration phase was i.i.d., but in general the rewards from a given arm are *not* i.i.d. due to adaptivity of action selections

#### Constructing confidence intervals

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Notation:

## Constructing confidence intervals

#### Let  $N_t^{(k)} = \sum_{k=1}^{k} 1_{\{a_k = k\}}$  be the number of times arm  $k$  is pulled before time *t* = *t*−1 ∑ *τ*=0

 $1_{\{a_{\tau}=k\}}$  be the number of times arm k is pulled before time t

#### Notation:

#### Constructing confidence intervals Let  $N_t^{(k)} = \sum_{k=1}^{k} 1_{\{a_k = k\}}$  be the number of times arm  $k$  is pulled before time *t* = *t*−1 ∑  $\tau=0$  $1_{\{a_{\tau}=k\}}$  be the number of times arm k is pulled before time t Let  $\hat{\mu}_t^{(k)} = \frac{1}{2\pi (k)} \sum_{k=1}^K \mathbb{1}_{\{a_r=k\}} r_\tau$  be the sample mean reward of arm k up to time *t* = 1  $N_t^{(k)}$ *t t*−1 ∑ *τ*=0  $1_{\{a_{\tau}=k\}} r_{\tau}$  be the sample mean reward of arm k up to time to Notation:



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So want Hoeffding to g

give us something like\n
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\left| \hat{\mu}_t^{(k)} - \mu \right| \le \sqrt{\frac{\ln(2/\delta)}{2N_t^{(k)}}} \text{ w/p } 1 - \delta
$$



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But this is generally FALSE (unless  $a_t$  chosen very simply, like exploration phase of ETC)



So want Hoeffding to give us something like

$$
\hat{\mu}_t^{(k)} - \mu
$$

- 
- 



### Constructing confidence intervals (cont'd)

The problem: Although  $r_{\tau} \mid a_{\tau} = k$  is an i.i.d. draw from  $\nu^{(k)},$  $= k$  is an i.i.d. draw from  $\nu^{(k)}$ 

## Constructing confidence intervals (cont'd)

(all  $arm$  indexing  $(k)$  now in superscripts; subscripts reserved for time index t) (*k*) *the problem: Although*  $r_\tau\mid a_\tau=k$  *is an i.i.d. draw from*  $\nu^{(k)}$ *,*  $\frac{q_{\text{all arm} \text{ ina} (k)} \text{ now in superscript a}}{s$  *alternal for time index*  $t$  $= k$  is an i.i.d. draw from  $\nu^{(k)}$ 



## Constructing confidence intervals (cont'd)

 $\hat{\mu}^{(k)}_t$  is the sample mean of a random number  $N^{(k)}_t$  of returns (all  $arm$  indexing  $(k)$  now in superscripts; subscripts reserved for time index t) (*k*) *the problem: Although*  $r_\tau\mid a_\tau=k$  *is an i.i.d. draw from*  $\nu^{(k)}$ *,*  $\frac{q_{\text{all arm} \text{ ina} (k)} \text{ now in superscript a}}{s$  *alternal for time index*  $t$  $= k$  is an i.i.d. draw from  $\nu^{(k)}$ 


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 $\hat{\mu}^{(k)}_t$  is the sample mean of a random number  $N^{(k)}_t$  of returns in general  $N_t^{(k)}$  will depend on those returns themselves *t*



(*k*)

 $\hat{\mu}^{(k)}_t$  is the sample mean of a random number  $N^{(k)}_t$  of returns in general  $N_t^{(k)}$  will depend on those returns themselves *t* (i.e., how often we select arm *k* depends on the historical returns of arm *k*)

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 $\widetilde{r}$ (*k*)  $\begin{array}{c} K(\mathcal{K})\ 0 \end{array}$  ,  $\widetilde{\mathcal{F}}$ (*k*)  $\begin{pmatrix} K \ 1 \end{pmatrix}, \widetilde{r}$ 

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- Solution: First, imagine an infinite sequence of *hypothetical* i.i.d. draws from  $ν<sup>(k)</sup>$ : (*k*)  $\begin{matrix} 2 \ 2 \end{matrix}$ ,  $\begin{matrix} \widetilde{r} \ \widetilde{r} \end{matrix}$ (*k*)  $\binom{K}{3}$ , ...



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Then we can think of every time we pull arm  $k$ , just pulling the next  $\tilde{r}$ 

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	- $\zeta^{(k)}$  off this list, *i*





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Then we can think of every time we pull arm  $k$ , just pulling the next  $\tilde{r}$ 

i.e.,  $r_{\tau} \mid a_{\tau} = k$  simply equal to  $\widetilde{r}_{\scriptscriptstyle{\cal N}^{(k)}}^{\scriptscriptstyle (K)}$ , and hence  $\overline{r}_{\scriptscriptstyle{\cal N}^{(k)}}^{\scriptscriptstyle (K)}$ 

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- Then we can think of every time we pull arm  $k$ , just pulling the next  $\widetilde{r}_{i}^{(k)}$  off this list, *i*<br>1. (*k*)  $N_{\tau}^{(k)}$ *μ* ̂ (*k*) *t* = 1  $N_t^{(k)}$ *t*  $N_t^{(k)}$   $-1$ ∑ *i*=0  $\widetilde{r}$ (*k*) *i*





### Constructing confidence intervals (cont'd) Recall: *μ* ̂ (*k*) *t* = 1  $N_t^{(k)}$ *t*  $N_t^{(k)} - 1$ ∑ *i*=0  $\widetilde{r}$ (*k*) *i*

### Constructing confidence intervals (cont'd) Recall: *μ* ̂ (*k*) *t* = 1  $N_t^{(k)}$ *t*  $N_t^{(k)} - 1$ ∑ *i*=0  $\widetilde{r}$ (*k*)  $\tilde{\mu}^{(K)}$  Now define:  $\tilde{\mu}$ *n* = 1 *n n*−1 ∑  $\widetilde{r}$ (*k*)  $\hat{\mu}^{(K)}$  (  $\Rightarrow \hat{\mu}$ ̂ (*k*) *t*

 $\tilde{\mu}_n^{(k)} = \frac{1}{L} \sum_{i} \tilde{r}_i^{(k)} \quad (\Rightarrow \hat{\mu}_t^{(k)} = \tilde{\mu}_{n(k)}^{(k)})$ *i*=0  $= \tilde{\mu}$ 



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*t*  $= \tilde{\mu}$ (*k*)

- $\tilde{\mu}_n^{(k)} = \frac{1}{L} \sum_{i} \tilde{r}_i^{(k)} \quad (\Rightarrow \hat{\mu}_t^{(k)} = \tilde{\mu}_{n(k)}^{(k)})$ *i*=0  $= \tilde{\mu}$
- Now Hoeffding applies to  $\tilde{\mu}_n^{(k)}$  because  $n$  fixed/nonrandom  $n^{(K)}$  because  $n$
- and we know  $\hat{\mu}_{t}^{(k)} = \tilde{\mu}_{n}^{(k)}$  for some  $n \leq t$  (but which one is *random*)  $n^{(K)}$  for some  $n \leq t$
- *i*=0  $n^{(K)}$  because  $n$
- and we know  $\hat{\mu}_{t}^{(k)} = \tilde{\mu}_{n}^{(k)}$  for some  $n \leq t$  (but which one is *random*)  $n^{(K)}$  for some  $n \leq t$ Can anyone suggest a strategy for getting a bound for  $|\,\hat{\mu}_{t}^{(k)}-\mu^{(k)}\,|$  ?  $\begin{array}{c} \hline \end{array}$

### Constructing confidence intervals (cont'd) Recall: *μ* ̂ (*k*) *t* = 1  $N_t^{(k)}$ *t*  $N_t^{(k)} - 1$ ∑ *i*=0  $\widetilde{r}$ (*k*)  $\tilde{\mu}^{(K)}$  Now define:  $\tilde{\mu}$  $\tilde{\mu}_n^{(k)} = \frac{1}{L} \sum_{i} \tilde{r}_i^{(k)} \quad (\Rightarrow \hat{\mu}_t^{(k)} = \tilde{\mu}_{n(k)}^{(k)})$ *n* = 1 *n n*−1 ∑  $\widetilde{r}$ (*k*)  $\hat{\mu}^{(K)}$  (  $\Rightarrow \hat{\mu}$ ̂ (*k*) *t*

Now Hoeffding applies to  $\tilde{\mu}_n^{(k)}$  because  $n$  fixed/nonrandom



*t*  $= \tilde{\mu}$ (*k*)





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\hat{\mu}_t^{(k)} = \tilde{\mu}_n^{(k)}
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 for so

Recall union bound in ETC analysis made Hoeffding hold simultaneously over  $k \leq K$ 

Now Hoeffding applies to  $\tilde{\mu}_n^{(k)}$  because  $n$  fixed/nonrandom  $n^{(K)}$  because  $n$ 

> $\tilde{u}^{(k)}_t = \tilde{\mu}^{(k)}_n$  for some  $n \leq t$  (but which one is *random*)  $n^{(K)}$  for some  $n \leq t$

Can anyone suggest a strategy for getting a bound for  $|\,\hat{\mu}_{t}^{(k)}-\mu^{(k)}\,|$  ?  $\begin{array}{c} \hline \end{array}$ 





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$$
\Rightarrow \mathbb{P}\left(\forall n \leq t, |\tilde{\mu}_n^{(k)} - \mu\right)
$$

- Now Hoeffding applies to  $\tilde{\mu}_n^{(k)}$  because  $n$  fixed/nonrandom  $n^{(K)}$  because  $n$ 
	- $\tilde{u}^{(k)}_t = \tilde{\mu}^{(k)}_n$  for some  $n \leq t$  (but which one is *random*)  $n^{(K)}$  for some  $n \leq t$
- Can anyone suggest a strategy for getting a bound for  $|\,\hat{\mu}_{t}^{(k)}-\mu^{(k)}\,|$  ?  $\begin{array}{c} \hline \end{array}$
- Recall union bound in ETC analysis made Hoeffding hold simultaneously over  $k \leq K$ 
	- Hoeffding + union bound over  $n \leq t$ :  $\left(\frac{k}{n} - \mu^{(k)}\right) \le \sqrt{\ln(2t/\delta)/2n} \ge 1 - \delta$

### Constructing confidence intervals (cont'd) Hoeffding + union bound over  $n \leq t$ :  $\Rightarrow P(\forall n \leq t, |\tilde{\mu}|)$  $\left| \mu^{(k)} - \mu^{(k)} \right| \leq \sqrt{\ln(2t/\delta)/2n} \big| \geq 1 - \delta$

### Constructing confidence intervals (cont'd) Hoeffding + union bound over  $n \leq t$ :  $\Rightarrow P(\forall n \leq t, |\tilde{\mu}|)$  $\left| \mu^{(k)} - \mu^{(k)} \right| \leq \sqrt{\ln(2t/\delta)/2n} \big| \geq 1 - \delta$ But since in particular  $N_t^{(k)} \leq t$ , this immediately implies  $t^{(K)} \leq t$ (*k*)  $|\leq \sqrt{\ln(2t/\delta)/2N_t^{(k)}}$  $\binom{r(k)}{t}$   $\geq 1-\delta$

$$
\Rightarrow \mathbb{P}\left(\forall n \leq t, \,|\tilde{\mu}_n^{(k)} - \mu\right)
$$

$$
\mathbb{P}\left(\left|\tilde{\mu}_{N_t^{(k)}}^{(k)} - \mu^{(k)}\right| \leq 1\right)
$$

And then since  $\tilde{\mu}_{_M(k)}^{(k)}=\hat{\mu}_{_L}^{(k)},$  we immediately get the kind of result we want:  $N_t^{(k)}$  $=$   $\hat{\mu}$ ̂ (*k*) *t* <sup>ℙ</sup> (|*<sup>μ</sup>* ̂  $\Rightarrow P(\forall n \leq t, |\tilde{\mu}|)$ But since in particular  $N_t^{(k)} \leq t$ , this immediately implies  $\mathbb{P}\setminus |\tilde{\mu}|$ (*k*)  $N_t^{(k)}$ 

### Constructing confidence intervals (cont'd) Hoeffding + union bound over  $n \leq t$ :  $\left| \mu^{(k)} - \mu^{(k)} \right| \leq \sqrt{\ln(2t/\delta)/2n} \big| \geq 1 - \delta$  $t^{(K)} \leq t$

$$
\left| \frac{d}{dt} \right| \leq \sqrt{\ln(2t/\delta)/2N_t^{(k)}} \geq 1 - \delta
$$

$$
-\mu^{(k)} \le \sqrt{\ln(2t/\delta)/2N_t^{(k)}} \ge 1 - \delta
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# Constructing confidence intervals (cont'd)

- Hoeffding + union bound over  $n \leq t$ :
	- $\left| \mu^{(k)} \mu^{(k)} \right| \leq \sqrt{\ln(2t/\delta)/2n} \big| \geq 1 \delta$
- But since in particular  $N_t^{(k)} \leq t$ , this immediately implies  $t^{(K)} \leq t$

$$
\left| \mu^{(k)} - \mu^{(k)} \right| \le \sqrt{\ln(2t/\delta)/2N_t^{(k)}} \ge 1 - \delta
$$

<u>Summary</u>: to deal with problem of non-i.i.d. rewards that enter into  $\hat{\mu}_{t}^{(\kappa)},$  we used rewards' *conditional* i.i.d. property along with a union bound to get Hoeffding bound that is wider by just a factor of  $t$  in the log term ̂ (*k*) *t* 21



$$
-\mu^{(k)} \le \sqrt{\ln(2t/\delta)/2N_t^{(k)}} \ge 1-\delta
$$

 $\int_{t}^{(k)} + \sqrt{\ln(2t/\delta)/2N_t^{(k)}}$ *t* ]

i.e., ̂  $\mathbb{R}$ *μ* ̂  $\hat{\mu}^{(k)} - \sqrt{\ln(2t/\delta)/2N_t^{(k)}}, \hat{\mu}^{(k)}$ ̂

So we have a valid  $(1 - \delta)$  confidence interval (CI) for  $\mu^{(\kappa)}$  at time t from last equation:  $(1 - \delta)$  confidence interval (CI) for  $\mu^{(k)}$  at time t

> $\mathbb{P}\left(\|\hat{\mu}_{t}^{(k)} - \mu^{(k)}\| \leq \sqrt{\ln(2t/\delta)/2N_t^{(k)}}\right) \geq 1-\delta,$  $\mu^{(k)}$  |  $\leq \sqrt{\ln(2t/\delta)/2N_t^{(k)}}$  $\left(\begin{array}{c} \mathbf{r}(\mathbf{k}) \\ t \end{array}\right) \geq 1 - \delta$

i.e.,  $\mathbb{P}\setminus\lvert\hat{\mu}\rvert$ ̂  $\mu^{(k)}$  |  $\leq \sqrt{\ln(2t/\delta)/2N_t^{(k)}}$  $\mathbb{R}$ *μ* ̂  $\hat{\mu}^{(k)} - \sqrt{\ln(2t/\delta)/2N_t^{(k)}}, \hat{\mu}^{(k)}$ ̂  $\int_{t}^{(k)} + \sqrt{\ln(2t/\delta)/2N_t^{(k)}}$ 

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$$
\binom{r(k)}{t} \geq 1 - \delta,
$$

*t* ]



 $\mathbb{P}\setminus\lvert\hat{\mu}\rvert$ ̂  $\mu^{(k)}$  |  $\leq \sqrt{\ln(2t/\delta)/2N_t^{(k)}}$ 

> $\int_{t}^{(k)} + \sqrt{\ln(2t/\delta)/2N_t^{(k)}}$ *t* ]

$$
\left.\begin{matrix}r(k)\\t\end{matrix}\right| \geq 1-\delta,
$$

i.e., 
$$
\hat{\mu}_t^{(k)} - \sqrt{\ln(2t/\delta)/2N_t^{(k)}}, \hat{\mu}_t^{(k)} + \sqrt{1}
$$

But analysis easier if CIs are *uniformly valid* over time *t* and arm *k*

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 $\int_{t}^{(k)} + \sqrt{\ln(2t/\delta)/2N_t^{(k)}}$ *t* ]

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\begin{pmatrix}r(k) \\ t\end{pmatrix} \geq 1 - \delta,
$$

i.e., 
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\hat{\mu}_t^{(k)} - \sqrt{\ln(2t/\delta)/2N_t^{(k)}}, \hat{\mu}_t^{(k)} + \sqrt{\frac{1}{2t}}
$$

By same argument as last two slides using a union bound over Hoeffding applied to all  $\tilde{\mu}_{n}^{(\kappa)}$  for , and noting that  $N_t^{(\kappa)} \leq T$  for all  $t < T$ , we get: (*k*) *n*  $n \leq T$ , and noting that  $N_t^{(k)}$  $T^{(K)}$   $\leq T$  for all  $t < T$ 

So we have a valid  $(1 - \delta)$  confidence interval (CI) for  $\mu^{(\kappa)}$  at time t from last equation:  $(1 - \delta)$  confidence interval (CI) for  $\mu^{(k)}$  at time t

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 $\mu^{(k)}$  |  $\leq \sqrt{\ln(2t/\delta)/2N_t^{(k)}}$  $\mathbb{P}\setminus\lvert\hat{\mu}\rvert$ ̂ Г

> $\int_{t}^{(k)} + \sqrt{\ln(2t/\delta)/2N_t^{(k)}}$ *t* ]

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\hat{\mu}_t^{(k)} - \sqrt{\ln(2t/\delta)/2N_t^{(k)}}, \ \hat{\mu}_t^{(k)} + \sqrt{\frac{1}{2t}}
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But analysis easier if CIs are *uniformly valid* over time *t* and arm *k*

Valid for any bandit algorithm! Of independent statistical interest for interpreting results

$$
\mathbb{P}\left(\forall t < T, \|\hat{\mu}_t^{(k)} - \mu^{(k)}\| \le \sqrt{\ln(2T/\delta)/2N_t^{(k)}}\right) \ge 1 - \delta
$$

So we have a valid  $(1 - \delta)$  confidence interval (CI) for  $\mu^{(\kappa)}$  at time t from last equation:  $(1 - \delta)$  confidence interval (CI) for  $\mu^{(k)}$  at time t

By same argument made in ETC analysis, union bound over *K* makes coverage uniform over *k*:  $\mathbb{P} \left( \forall k \leq K, t < T, |\hat{\mu}| \right)$ ̂  $\mu^{(k)} - \mu^{(k)}$  $\leq \sqrt{\ln(2TK/\delta)/2N_t^{(k)}}$ 

 $\mu^{(k)}$  |  $\leq \sqrt{\ln(2t/\delta)/2N_t^{(k)}}$  $\mathbb{P}\setminus\lvert\hat{\mu}\rvert$ ̂ Г

So we have a valid  $(1 - \delta)$  confidence interval (CI) for  $\mu^{(\kappa)}$  at time t from last equation:  $(1 - \delta)$  confidence interval (CI) for  $\mu^{(k)}$  at time t

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\hat{\mu}_t^{(k)} - \sqrt{\ln(2t/\delta)/2N_t^{(k)}}, \hat{\mu}_t^{(k)} + \sqrt{\frac{1}{2t}}
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But analysis easier if CIs are *uniformly valid* over time *t* and arm *k*

$$
k\Big|_{22} \le \sqrt{\ln(2TK/\delta)/2N_t^{(k)}} \Big| \ge 1 - \delta
$$





$$
\mathbb{P}\left(\forall t < T, \|\hat{\mu}_t^{(k)} - \mu^{(k)}\| \le \sqrt{\ln(2T/\delta)/2N_t^{(k)}}\right) \ge 1 - \delta
$$



- Feedback from last lecture
- Recap
- Regret analysis of ETC
	- -greedy algorithm *ε*
- Confidence intervals for the arms
	- Upper Confidence Bound (UCB) algorithm

### Upper Confidence Bound (UCB) algorithm  $For t = 0, ..., T - 1:$ Choose the arm with the highest upper confidence bound, i.e.,  $a_t = \arg \max_{t \in [1]}$ *k*∈{1,…,*K*} *μ* ̂  $\int_{t}^{(k)} + \sqrt{\ln(2TK/\delta)/2N_t^{(k)}}$

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 $For t = 0, ..., T - 1:$  $a_t = \arg \max_{t \in [1]}$ *k*∈{1,…,*K*} *μ*  $\hat{\mu}$ ̂ (2) *t*  $\hat{\mu}$ ̂  $\int_{t}^{(2)} + \sqrt{\ln(2TK/\delta)/2N_t^{(2)}}$  $\hat{\mu}$ ̂  $\hat{\mu}$ ̂ (1) *t*  $\hat{\mu}$ ̂  $\sqrt{\ln(2TK/\delta)/2N_t^{(1)}}$  $\hat{\mu}$ ̂  $\sqrt{\ln(2TK/\delta)/2N_t^{(1)}}$ *μ*(1)

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 $\hat{\mu}$ ̂ (3) *t*  $\hat{\mu}$ ̂  $\sqrt{\ln(2TK/\delta)/2N_t^{(3)}}$  $\hat{\mu}$ ̂  $\sqrt{\ln(2TK/\delta)/2N_t^{(3)}}$ *μ*(3)

*μ*(2)

(we can't see the  $\mu^{(k)}$ )









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 $\int_{t}^{(2)} - \sqrt{\ln(2TK/\delta)/2N_t^{(2)}}$ 

(we can't see the  $\mu^{(k)}$ )







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Optimism in the face of uncertainty is an important principle in RL It basically says to give each arm the benefit of the doubt, and basically act as if that arm is as good as it could plausibly be in choosing an action



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Since each upper bound is 
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Note that the exploration here is *adaptive*, i.e., focused on most promising arms







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## Summary:

Feedback: [bit.ly/3RHtlxy](http://bit.ly/3RHtlxy)

- ETC and  $\varepsilon$ -greedy, achieve sublinear regret  $O$
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## Attendance: [bit.ly/3RcTC9T](http://bit.ly/3RcTC9T)



 $\tilde{O}(T^{2/3})$ 

• Hoeffding can be used to provide (uniform) bounds on the arm means • UCB algorithm follows "optimism in the face of uncertainty" principle